

# High Dimensional Multivariate Realized Volatility Measures

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# The general setup and motivation

- $p$  assets,  $p$  large;
- Frictionless log-prices: a  $p$ -dimensional  $Itô$  process with factors

$$dP_t^* = b \cdot \sigma_{ft} dB_t^F + \sigma_{\epsilon t} dB_t^I \quad (1)$$

- Object of interest: the integrated covolatility matrix of asset prices

$$\Sigma_T = b \cdot \int_0^T (\sigma_{ft} \sigma'_{ft}) dt \cdot b' + \int_0^T (\sigma_{\epsilon t} \sigma'_{\epsilon t}) dt \quad (2)$$

- $\Sigma_T$ : Important for asset pricing, portfolio allocation and risk management;
  - Optimal portfolio allocation: function of  $\Sigma_T^{-1}$ ;
  - Risk of a portfolio: related to  $\Sigma_T$ .
- The portfolio allocation or the risk management can involve a large number of assets;
- Need to have an accurate and well conditioned estimator for a large number of assets.

# Related literature

- Estimators of the covolatility for a small number of assets: Hayashi et al. (2005); Barndorff-Nielsen et al.(2008); Christensen et al. (2010);
- High dimensional cases: Nerlov and Diebold (1989); Shephard et al. (2006); Fan et al. (2008); Wang et al.(2010); Lunde et al. (2011); Tao et al. (2011) and Bannouh et al. (2012); Ait-Sahalia et al. (2016)
- For most of these papers: sparsity; factor structure of returns; observable factors;
- Problem: no consensus on factors to use; misspecification and missing factors problems;
- **This paper**: factor structure of returns; unobserved factors; **microstructure noise; factor structure on microstructure noise**;
- *Advantage of a factor model*: a semi-definite positive estimator; Estimator is invertible under weak conditions; Reduction of the number of parameters to estimate.

# Key findings

- When the number of assets is large, a factor structure in returns and in microstructure noise improve the estimation of:
  - The integrated covolatility matrix;
  - The correlation matrix;
  - The inverse of the integrated covolatility matrix.

# Outline

- 1 Introduction
- 2 Theoretical framework and methodology
- 3 Sampling properties and simulations
- 4 Empirical studies
- 5 Concluding remarks

# The setup

- $\Delta$ : frequency of the data

- $dP_t^* = b \cdot dF_t + dE_t \implies \int_{t-\Delta}^t dP_s^* = b \cdot \int_{t-\Delta}^t dF_s + \int_{t-\Delta}^t dE_s$

- Notations:

- $r_t^* \equiv r_{t,\Delta}^* \equiv \int_{t-\Delta}^t dP_s^*$ ;

- $f_t \equiv f_{t,\Delta} \equiv \int_{t-\Delta}^t dF_s$ ;

- $\epsilon_t \equiv \epsilon_{t,\Delta} \equiv \int_{t-\Delta}^t dE_s$

- Matricial representation:

$$r_t^* = b f_t + \epsilon_t \quad (3)$$

- $f_{kt} \perp f_{lt}, \forall k \neq l$ ;

- $f_{kt} \perp \epsilon_{it}$ ;

- $\epsilon_{it} \perp \epsilon_{jt}, \forall i \neq j$ ;

- $E(\epsilon_{it} | I_{t-\Delta}) = 0$

# The setup

- The covolatility matrix

$$\Sigma = b\Sigma_f b' + \Sigma_\epsilon \quad (4)$$

- Parameters of interest:  $\Sigma_f$ ,  $\Sigma_\epsilon$ , and  $b$
- Observed log-prices are noisy

$$P_t = P_t^* + u_t \quad (5)$$

- Factor representation of the microstructure noise:

$$u_t = cg_t + \eta_t \quad (6)$$

- $\eta_t \perp g_t, f_t, \epsilon_t, P_t^*$ ;
- $g_t \perp f_t, \epsilon_t, P_t^*$ ;
- $g_t$  and  $\eta_t$ : independent across time;
- $c$  is a  $p \times K'$  vector of noise loadings;
- $g_t$  is  $K' \times 1$ : captures the cross-sectional dependence in the measurement errors

## Evidence of a factor structure in microstructure noise

- 384 stocks traded on *S&P500*;
- Six random trading days;
- For each day  $d$ , estimation of the covolatility of the microstructure noise:

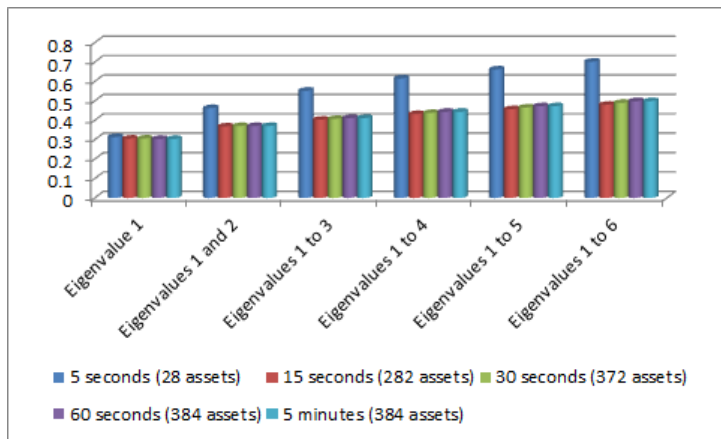
$$\hat{\Sigma}_u^d = \frac{\sum_{i=1}^n r_{j/n} r'_{j/n}}{2n}$$

- Run the spectral decomposition of  $\hat{\Sigma}_u^d$ ,  $d = 1, \dots, 6$  to get eigenvalues;
- For different sampling frequency, compute the share of the total variability of the microstructure noise explained 1, 2, ..., 6 factors.



## Evidence of a factor structure in microstructure noise

Figure: Ratio of largest eigenvalues relative to the total variation

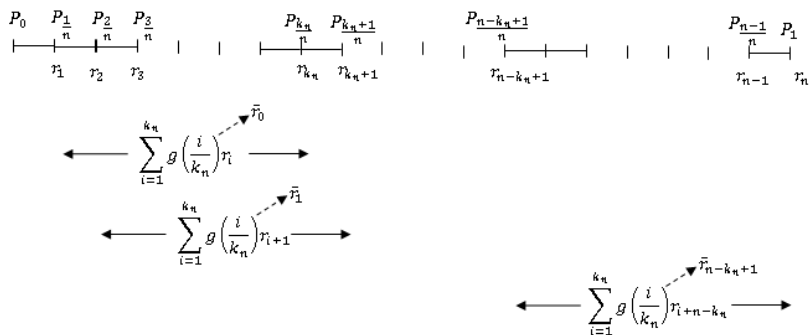


# Estimation: recall on the Pre-averaging approach

- Observed returns are also noisy

$$r_{it} = r_{it}^* + (u_{it} - u_{it-\Delta})$$

Figure: The pre-averaging approach



- The probability order of the microstructure noise part moves from  $O_p(1)$  to  $O_p\left(\frac{1}{\sqrt{k_n}}\right)$ .

# Estimators based on the Pre-averaging approach

- Pre-averaging estimator of the integrated volatility (Jacod et al. (2009))

$$PRV(r) = \frac{1}{\theta\psi_2\sqrt{n}} \sum_{j=0}^{n-k_n+1} (\bar{r}_j)^2 - \frac{\psi_1}{2\theta^2\psi_2n} \sum_{i=1}^n (r_i)^2 \quad (7)$$

- Pre-averaging estimator of the integrated covolatility (Christensen et al. (2010))

$$MRC[r]_n = \frac{n}{(n-k_n+2)} \frac{1}{\psi_2k_n} \sum_{i=0}^{n-k_n+1} \bar{r}_i \bar{r}'_i - \frac{\psi_1^{k_n}}{2n\theta^2\psi_2^{k_n}} \sum_{i=1}^n r_i r'_i \quad (8)$$

$\theta$ ,  $\psi_1$  and  $\psi_2$ : setting parameters;  $k_n$ : number of averaged return per block;  $\bar{r}_i$  the  $i^{th}$  averaged return,  $r_i$  the  $i^{th}$  usual return.

# The different steps: an overview

- Estimation of rotations of the factor components using the factor analysis;
- Estimation of the integrated volatility of factors;
- Estimation of factor loadings;
- Estimation of idiosyncratic components;
- Estimation of the integrated volatility of idiosyncratic components;
- Deduction of  $\hat{\Sigma}$  by a natural plug-in.

# Estimation of a rotation of factors

- $\Delta$ : The observation frequency assumed to be constant over time and over assets, and closed to zero;
- Estimation of the rotated factor  $\tilde{f}$ : minimization of the scaled sum of squared values of the idiosyncratic component:

$$\left\{ \begin{array}{l} \text{Min}_{f_{j\Delta}, b} \frac{1}{p} \sum_{j=1}^{\lfloor 1/\Delta \rfloor} (r_{j\Delta}^* - bf_{j\Delta})'(r_{j\Delta}^* - bf_{j\Delta}) \\ \text{s.t. } \frac{1}{p} b'b = I_K \end{array} \right. \quad (9)$$

- The solution of the minimization problem

$$\hat{f}_{j\Delta} = \frac{1}{p} W' r_{j\Delta}^*, \quad \forall j = 1, \dots, \lfloor 1/\Delta \rfloor$$

- $W$  is the matrix of ordered eigenvectors of  $\sum_{j=1}^{\lfloor 1/\Delta \rfloor} r_{j\Delta}^* r_{j\Delta}^{*'}.$

# Estimation of a rotation of factors

- The continuous time solution: by taking  $\Delta \rightarrow 0$

$$\hat{f}_t = \frac{1}{p} W' r_t^*, \quad \forall t > 0 \quad (10)$$

- Columns of  $W$  correspond to ordered eigenvectors of  $\Sigma$ ;
- $W$  and  $r^*$  are latent: unfeasible estimator;
- As feasible estimator:

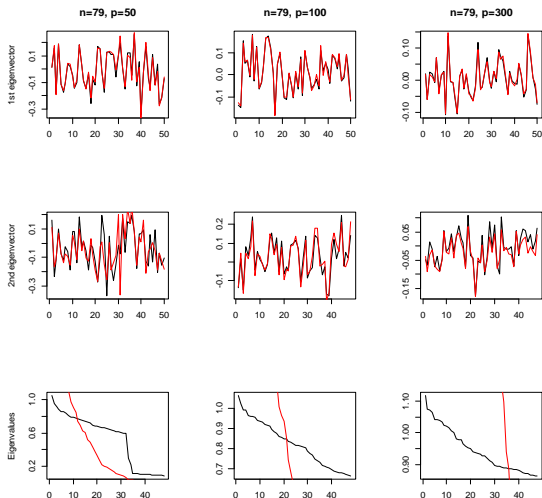
$$\hat{f}_t = \frac{1}{p} \hat{W}' r_t, \quad \forall t > 0 \quad (11)$$

- $\hat{W}$ : the matrix of  $K$  ordered eigenvectors of  $MRC$  (or another consistent estimator in the small dimensional framework as the Kernel estimator

$$MRker \text{ defined by } MRker(r) = \sum_{h=-n}^n k\left(\frac{h}{H+1}\right) \sum_{j=h+1}^n r_j r'_{j-h},$$

- $r_t$  the series of the  $p \times 1$  vector of observed returns.

## Why to use eigenvectors of Kernel or Pre-averaging covolatility estimators?

Figure: Spectral decomposition of  $MRker$ 

# Estimation of parameters of interest

- $\hat{f}_t$  can be decomposed as

$$\begin{aligned}
 \hat{f}_{kt} &= \frac{1}{p} W'_k r_t^* + \frac{1}{p} W'_k (u_t - u_{t-\Delta}) + \frac{1}{p} W_k^{\epsilon'} r_t^* + \frac{1}{p} W_k^{\epsilon'} (u_t - u_{t-\Delta}) \\
 &\approx \frac{1}{p} W'_k b f_t + \frac{1}{p} W'_k \epsilon_t + \frac{1}{p} W'_k (u_t - u_{t-\Delta}) \\
 &\approx \frac{1}{p} W'_k b f_t + \frac{1}{p} W'_k (u_t - u_{t-\Delta}) \\
 &\approx \hat{f}_{kt} + \frac{1}{p} W'_k (u_t - u_{t-\Delta}) \\
 &\approx \tilde{f}_{kt} + \frac{1}{p} W'_k c (g_t - g_{t-\Delta}) + O_p(p^{-1/2} n^{-1/2})
 \end{aligned}$$

- $\hat{f}$ : consistently estimate a rotation of the latent factor  $f$  contaminated by microstructure noises.



# Estimation of parameters of interest

- **Remark:** Let  $H$  be a  $K \times K$  orthogonal matrix ( $H'H = I_K$ ). Then

$$\begin{aligned}
 r_t^* &= bf_t + \epsilon_t \\
 &= bI_K f_t + \epsilon_t \\
 &= bH'H f_t + \epsilon_t \\
 &= \tilde{b}\tilde{f}_t + \epsilon_t, \quad \text{with } \tilde{b} = bH' \text{ and } \tilde{f}_t = Hf_t
 \end{aligned}$$

- *Conclusion:*  $\Sigma$  is unchanged by the same rotation of the factors and loadings;

## Estimator of the integrated volatility of $\tilde{f}$

- $\hat{f}_{kt} = \tilde{f}_{kt} + \frac{1}{p} W_k' c(g_t - g_{t-\Delta}) + O_p(p^{-1/2}n^{-1/2})$
- $\int_0^1 \widehat{\sigma_{fku}^2} du = PRV(\hat{f}_k)$
- $PRV()$  is the pre-averaging estimator defined by (8).

# Estimation of parameters of interest

- *Notation:*  $[X, Y]$  denotes the integrated covolatility between  $X$  and  $Y$ ;

$$\begin{aligned}
 [r_i^*, \tilde{f}_k] &= [\tilde{b}_{ik} \tilde{f}_k + \epsilon_i, \tilde{f}_k] \\
 &= \tilde{b}_{ik} [\tilde{f}_k, \tilde{f}_k] + [\epsilon_i, \tilde{f}_k] \\
 &= \tilde{b}_{ik} \int_0^1 \sigma_{\tilde{f}_{kt}}^2 dt
 \end{aligned}$$

## Estimator of the loadings

- $[r_i^*, \tilde{f}_k] = \tilde{b}_{ik} \int_0^1 \sigma_{\tilde{f}_{kt}}^2 dt \implies \hat{b}_{ik} = \frac{MRC(r_i, \hat{f}_k)}{PRV(\hat{f}_k)}$
- $MRC$  is the Modulated realized covariance defined by (10).

# Estimation of parameters of interest

For  $n$  and  $p$  are sufficiently large

$$\begin{aligned}
 \hat{\epsilon}_{it} &= r_{it} - \sum_{k=1}^K \hat{b}_{ik} \hat{f}_{kt} \\
 &\approx r_{it} - \sum_{k=1}^K \tilde{b}_{ik} \tilde{f}_{kt} - \frac{1}{p} \sum_{k=1}^K \tilde{b}_{ik} W_k' c(g_t - g_{t-\Delta}) \\
 &\approx \epsilon_{it} + (u_t - u_{t-\Delta}) - \frac{1}{p} \sum_{k=1}^K \tilde{b}_{ik} W_k' c(g_t - g_{t-\Delta})
 \end{aligned}$$

$\hat{\epsilon}_{it}$  consistently estimates  $\epsilon_{it}$  plus a microstructure noise term;

Estimator of the integrated volatility of the idiosyncratic components

- $$\int_0^1 \widehat{\sigma_{\epsilon i u}^2} du = PRV(\hat{\epsilon}_i)$$

# Estimator of the covolatility matrix

- $\hat{\Sigma}$  is obtained by plug in the estimators of  $\tilde{b}$ ,  $\int_0^1 \sigma_{f_{ku}}^2 du$  for  $k = 1, \dots, K$  and  $\int_0^1 \sigma_{\varepsilon_{iu}}^2 du$ , for  $i = 1, \dots, p$  in (16)

## The covolatility matrix estimator

$\forall i, j = 1, \dots, p$

- $\hat{\Sigma}_{ij} = \sum_{k=1}^K \frac{MRC(r_i, \hat{f}_k) \cdot MRC(r_j, \hat{f}_k)}{PRV(\hat{f}_k)}$
- $\hat{\Sigma}_{ii} = \sum_{k=1}^K \frac{MRC(r_i, \hat{f}_k)^2}{PRV(\hat{f}_k)} + PRV(\hat{\varepsilon}_i)$

- $PRV$  and  $MRC$  are respectively the uni dimensional and the multidimensional pre-averaging estimators.

# Competitors

- The Multidimensional Realized kernel

$$MRker(r) = \sum_{h=-n}^n k\left(\frac{h}{H+1}\right)\Gamma_h, \quad (12)$$

$$\Gamma_h = \sum_{j=h+1}^n r_j r'_{j-h}, \text{ for } h > 0; \Gamma_h = \Gamma'_{-h}, \text{ for } h < 0.$$

- The adjusted Modulated Realized Covariance

$$MRC[r]_n^\delta = \frac{n}{(n - k_n + 2)} \frac{1}{\psi_2 k_n} \sum_{i=0}^{k_n} (\bar{r}_i) (\bar{r}_i)' \quad (13)$$

- The composite Kernel  $\hat{\Sigma}_{comp}$ : uni-realized Kernel for diagonal elements, bivariate realized kernel for off-diagonal terms and regularization.

# Rates of convergence

## Lemma

Under assumptions of the model and the Frobenius norm, we have the following results

- $\left| \hat{\Sigma}_{kk}^{\tilde{f}} - \Sigma_{kk}^{\tilde{f}} \right| = O_p \left( n^{-1/4} \right), \forall k = 1, \dots, K$
- $\left\| \hat{b}_k - b_k \right\|_F = O_p \left( p^{1/2} n^{-1/4} \right), \forall k = 1, \dots, K$
- $\left\| \hat{\Sigma}^{\epsilon} - \Sigma^{\epsilon} \right\|_F = O_p \left( p^{1/2} n^{-1/4} \right)$

## Theorem

Under assumptions of the model, we have

- $\left\| \hat{\Sigma} - \Sigma \right\|_F = O_p \left( p n^{-1/4} \right)$
- $\left\| MRC^{\delta} - \Sigma \right\|_F = O_p \left( p n^{-1/5} \right)$
- $\left\| MRker - \Sigma \right\|_F = O_p \left( p n^{-1/5} \right)$
- $\left\| \hat{\Sigma}_{comp} - \Sigma \right\|_F = O_p \left( \sqrt{p(p-1)} n^{-1/5} \right)$

# Simulation design

- The factor component:

$$f_{kt} = \sigma_{f_{kt}} dB_t^F, k = 1, 2$$

$B_t^F$  a brownian motion and  $\sigma_{f_{kt}} \sim GARCH$  diffusion model as in Torben G. Andersen et al. (2006);

- The idiosyncratic error term:

$$\varepsilon_{it} = \sigma_{\varepsilon_{it}} dB_{it}^I$$

- For  $1 \leq i \leq p/3$ ,  $\sigma_{\varepsilon_{it}} \sim Nelson GARCH$  diffusion limit model as in Bandorff-Nielsen and Shephard (2002);
- For  $p/3 < i \leq 2p/3$ ,  $\sigma_{\varepsilon_{it}} \sim OU$  model as in Bandorff-Nielsen and Shephard (2002);
- For  $2p/3 < i \leq p$ ,  $\sigma_{\varepsilon_{it}} \sim GARCH$  diffusion model as in Torben G. Andersen et al. (2006).
- The microstructure noise: one factor.

## Simulation results: the synchronous case

**Table:** Scaled Frobenius norm  $\|A\|_F = \sqrt{\frac{1}{p} \left( \sum_{i=1}^p \sum_{j=1}^p |a_{ij}^2| \right)}$  of the estimation errors for the factorial estimator, the composite kernel, the multivariate Kernel and the adjusted modulated realized covariance: data are synchronous and low noise ( $\xi^2 = 0.001$ );  $n = 78$

Number of assets: p=50						
	Covariance	Correlation	Inverse	PSD	Diag	Off-Diag
$\hat{\Sigma}$	<b>1.206</b> (0.283)	<b>0.921</b> (0.190)	<b>2.271</b> (5.389)	<b>1</b>	<b>5.183</b>	<b>75.007</b>
$\hat{\Sigma}_{comp}$	1.541 (0.292)	1.055 (0.131)	3.170 (0.197)	1	11.940	119.563
MRker	1.587 (0.279)	1.128 (0.129)	270.884 (74.929)	1	6.281	134.326
$MRC^\delta$	1.560 (0.279)	1.116 (0.132)	147.141 (395.209)	1	6.025	128.28
Number of assets: p=100						
$\hat{\Sigma}$	<b>2.310</b> (0.485)	<b>1.221</b> (0.205)	<b>2.895</b> (34.790)	<b>1</b>	<b>18.498</b>	<b>538.322</b>
$\hat{\Sigma}_{comp}$	2.730 (0.944)	1.418 (0.186)	3.835 (0.213)	1	62.401	910.124
MRker	2.778 (0.923)	1.498 (0.184)	NA NA	0	27.7020	963.417
$MRC^\delta$	2.698 (0.864)	1.480 (0.190)	NA NA	0	24.707	891.115



## Simulation results: the asynchronous case

**Table:** Scaled Frobenius norm  $\|A\|_F = \sqrt{\frac{1}{p} \left( \sum_{i=1}^p \sum_{j=1}^p |a_{ij}^2| \right)}$  of the estimation errors for the factorial estimator, the composite kernel, the multivariate Kernel and the adjusted modulated realized covariance: data are asynchronous and low noise ( $\xi^2 = 0.001$ )

Number of assets: p=50						
	Covariance	Correlation	Inverse	PSD	Diag	Off-Diag
$\hat{\Sigma}$	<b>4.559</b> (0.317)	<b>1.648</b> (0.120)	<b>3.424</b> (0.196)	<b>1</b>	<b>94.875</b>	<b>963.622</b>
$\hat{\Sigma}_{comp}$	4.790 (0.279)	1.629 (0.069)	3.250 (0.141)	1	90.197	1075.323
MRker	4.906 (0.279)	1.710 (0.079)	229.652 (111.995)	1	89.754	1141.481
$MRC^\delta$	4.829 (0.278)	1.700 (0.079)	128.104 (62.922)	1	87.405	1110.009
Number of assets: p=100						
$\hat{\Sigma}$	<b>6.329</b> (0.448)	<b>2.131</b> (0.141)	<b>3.426</b> (0.595)	<b>1</b>	<b>170.842</b>	<b>3912.406</b>
$\hat{\Sigma}_{comp}$	6.908 (0.400)	2.347 (0.083)	3.558 (0.133)	1	146.474	4678.125
MRker	7.233 (0.412)	2.511 (0.107)	NA NA	0	165.132	5061.262
$MRC^\delta$	7.002 (0.422)	2.416 (0.109)	NA NA	0	152.773	4752.494

# Empirical studies

- **Data:** Five years of intraday data, collected from the TAQ database of the Wharton Research Data Services (January 2007-December 2011);
- The choice of the number of factors as in Bai et al. (2002): minimizing the criteria

$$IC = \log \left[ \frac{1}{p} \sum_{j=1}^{\lfloor 1/\Delta \rfloor} (r_{j\Delta} - bf_{j\Delta})'(r_{j\Delta} - bf_{j\Delta}) \right] + Kg(p, \lfloor 1/\Delta \rfloor) \quad (14)$$

- Performance evaluation
  - Resolve the following classic Markowitz portfolio problem with a gross exposure constraint as in Fan et al. (2008)

$$\begin{cases} \text{Min} & \hat{w}'_{t+1} \hat{\Sigma}_t w_{t+1} \\ \text{s.t} & \hat{w}'_{t+1} \mathbf{1} = 1 \quad \text{and} \quad \sum_{i=1}^p |w_{i,t+1}| \leq 1 + 2s \end{cases} \quad (15)$$

- The performance criteria

$$\hat{w}'_{t+1} RCov_{t+1} \hat{w}_{t+1} \quad (16)$$

where  $RCov_{t+1}$  is the out of sample 5-minutes realized covariance

## Empirical studies

**Table:** Performance evaluation of estimators of the covolatility matrix: Average out of sample risk under gross exposure constraints; liquid assets (traded at least 195 times)

	s=0	s=0.25	s=0.5	s=1
$\hat{\Sigma}$	<b>0.495</b>	<b>0.245</b>	<b>0.240</b>	<b>0.241</b>
$\hat{\Sigma}_{comp}$	0.520 (0.004)	0.301 (0.003)	0.325 (0.001)	0.326 (0.001)
<i>MRKer</i>	0.545 (0.04)	0.278 (0.09)	0.263 (0.080)	0.258 (0.007)
<i>MRC<math>\delta</math></i>	0.559 (0.014)	0.362 (0.014)	0.343 (0.011)	0.319 (0.003)
<i>EqualWeight</i>	0.636 (6E-05)	0.636 (5E-09)	0.636 (3E-09)	0.636 (2E-09)

# Conclusion

- We use a statistical factor model to estimate a high dimensional covolatility matrix;
- We provide evidence of a factor structure the microstructure noise;
- A factor structure in returns and in microstructure noise improve the estimation of the integrated covolatility matrix, the correlation matrix, the inverse of the integrated covolatility matrix.;
- This result is proved analytically, confirmed through a simulation exercise and an empirical study.

THANK YOU